

Understanding Small-Sample Statistics for Experimental Analysis

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Overview

Deficiencies in Current Research

- 1 Generally understood that small samples are a problem
- 2 Not generally understood why they are a problem
- 3 Coordinate on certain favorite techniques
- 4 Favorite techniques are not always appropriate

Overview

Presentation Overview

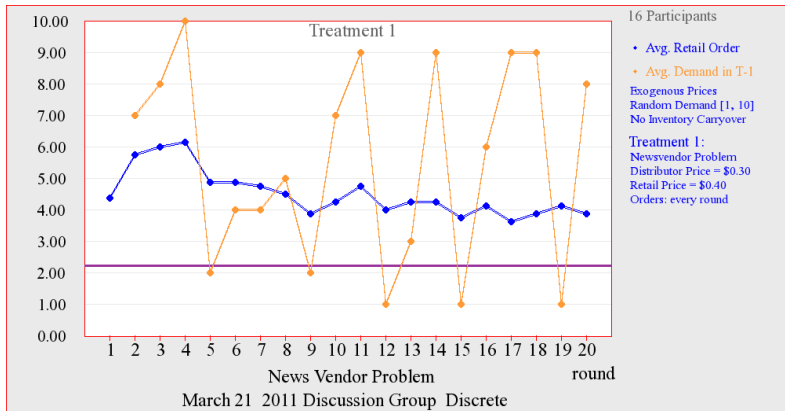
- 1 Intuition behind problem and solutions
- 2 Compare alternative tests
- 3 Application to Newsvendor Problem

Outline

- 1 Newsvender Problem
- 2 Normal Theory
- 3 Fisher Permutation
- 4 Wilcoxon Signed-Rank
- 5 Sign Test
- 6 Comments

Newsvendor Problem

Figure: Newsvendor Problem Graph



Newsvendor Problem Data

Average Quantity Supplied by Participant

$$q_1 = 4.00 \quad q_5 = 3.85$$

$$q_2 = 4.20 \quad q_6 = 4.65$$

$$q_3 = 5.75 \quad q_7 = 3.60$$

$$q_4 = 5.60 \quad q_8 = 4.35$$

Data Properties

- independent ✓
- identically distributed ✓
- small sample, $n = 8$ ✓

Research Question

Summary of Current Data

$$\bar{q} = \frac{1}{n} \sum_{i=1}^n q_i = 4.5 \quad s^2 = \frac{1}{n-1} \sum_{i=1}^n (q_i - \bar{q})^2 = 0.792$$

Inferential Question

If **another group** of subjects played the **same game** tomorrow, do these data suggest tomorrow's average quantity supplied would exceed 4 units?

Simplifying Trick

Transformation

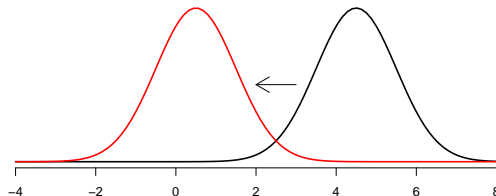
Let $x_i = q_i - 4$ as follows:

$$q_1 = 4.00 \quad q_5 = 3.85 \quad \rightarrow \quad x_1 = 0.00 \quad x_5 = -0.15$$

$$q_2 = 4.20 \quad q_6 = 4.65 \quad \rightarrow \quad x_2 = 0.20 \quad x_6 = 0.65$$

$$q_3 = 5.75 \quad q_7 = 3.60 \quad \rightarrow \quad x_3 = 1.75 \quad x_7 = -0.40$$

$$q_4 = 5.60 \quad q_8 = 4.35 \quad \rightarrow \quad x_4 = 1.60 \quad x_8 = 0.35$$



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Test Statistic

Familiar Normal-Theory Test Statistic

$$T_{STT} = \frac{\bar{x}}{\sqrt{V[\bar{x}]}} = \frac{\bar{x}}{s/\sqrt{n}}$$

$$\bar{x} = 0.5$$

$$s = 0.792$$

$$n = 8$$

$$T_{STT}^* = 1.786$$

Null Distribution

Normal Distribution, Any Sample Size

If we assume the sample observations follow a normal distribution so that $x_1, \dots, x_n \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$ then $T_{STT} \sim t_{n-1}$.

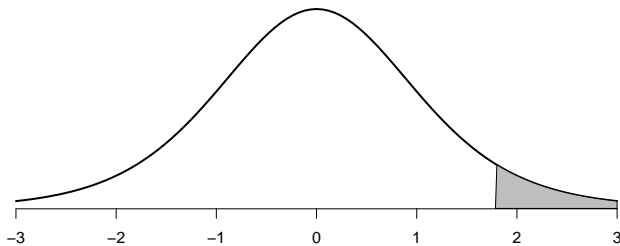
Any Distribution, Large Sample Size

As the sample size becomes large so that $n \rightarrow \infty$, for any mean-zero distribution of sample observations, $T_{STT} \xrightarrow{d} t_{n-1} \xrightarrow{d} N(0, 1)$.

Null Distribution

Calculating the P-Value

$$\begin{aligned} \text{p-value} &= P(T_{STT} \geq T_{STT}^*) \\ &= P(t_7 \geq 1.786) \\ &= 0.0586 \end{aligned}$$



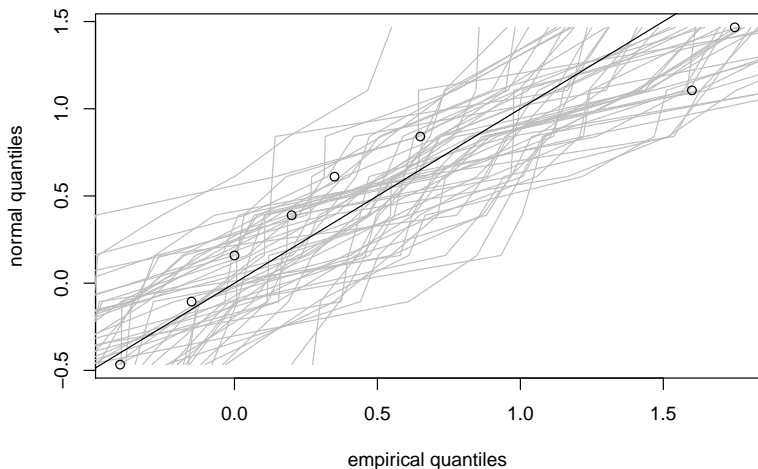
Null Distribution

Non-Normal Distribution, Small Sample Size

- 1 When **skewness** or **kurtosis** are severe, results of the t test can be misleading.
- 2 The sample size n needed to achieve good results increases as the sample distribution becomes less normal.
- 3 The t test is rather **robust for validity**, but not generally **robust for efficiency**.

Checking Normality Assumption

Figure: Normal Probability (Probit) Plot



Options when Normality Violated

- 1 More robust test statistics
- 2 Transformations of sample data
- 3 Nonparametric test statistics ✓

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Test Statistic

Fisher Permutation Test Statistic

$$T_{FPT} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x} \propto T_{STT}$$

$$T_{FPT}^* = \bar{x} = 0.5$$

Deriving the Null Distribution

Symmetric Distribution

If the sampling distribution of the observations is **symmetric around zero**, then any observation in the sample was equally likely to have been drawn with either a **positive** or **negative** sign.

Sign Assignments

The particular signs observed in the sample are only one out of $2^n = 256$ equally likely ways that signs could have attached to the observed data.

Test Statistic Values

Each vector of signs yields a **test statistic** with a $1/256$ chance of being observed. The full set of 256 test statistic values **is the null distribution**.

Deriving the Null Distribution

Example Construction of Null Distribution for $n = 3$

Consider only the last three observations: 0.65, -0.40, 0.35.

x_6	x_7	x_8	\rightarrow	\bar{x}
0.65	0.4	0.35	\rightarrow	0.466
0.65	0.4	-0.35	\rightarrow	0.233
0.65	-0.4	0.35	\rightarrow	0.200
0.65	-0.4	-0.35	\rightarrow	-0.033
-0.65	0.4	0.35	\rightarrow	0.033
-0.65	0.4	-0.35	\rightarrow	-0.200
-0.65	-0.4	0.35	\rightarrow	-0.233
-0.65	-0.4	-0.35	\rightarrow	-0.466

Null Distribution

Calculating the P-Value

$$\begin{aligned} \text{p-value} &= P(T_{FPT} \geq T_{FPT}^*) \\ &= \#\{T_{FPT} \geq T_{FPT}^*\} / 2^n \\ &= 18/256 \\ &= 0.0703125 \end{aligned}$$

Comments

- 1 No **closed-form**, **recursive**, or **approximate** definition of null distribution
- 2 Computation cost **exponential** in sample size
- 3 **Bootstrapping** null distribution can reduce computation cost for large samples
- 4 **Asymptotically equivalent** to the usual t test

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Test Statistic

Signed-Rank Definition

Define SR_i as the **signed-rank** of observation x_i in sample x :

x_i	0	0.2	1.75	1.6	-0.15	0.65	-0.4	0.35
SR_i	1	3	8	7	-2	6	-5	4

Wilcoxon Signed-Rank Test Statistic

$$T_{WSR} = \sum_{i=1}^n SR_i$$

$$T_{WSR}^* = 1 + 3 + 8 + 7 - 2 + 6 - 5 + 4 = 22$$

Deriving the Null Distribution

Symmetric Distribution

If the sampling distribution of the observations is **symmetric around zero**, then any observation in the sample was equally likely to have been drawn with either a **positive** or **negative** sign.

Sign Assignments

The particular signs observed in the sample are only one out of $2^n = 256$ equally likely ways that signs could have attached to the observed data.

Test Statistic Values

Each vector of signs yields a **test statistic** with a $1/256$ chance of being observed. The full set of 256 test statistic values **is the null distribution**.

Deriving the Null Distribution

Example Construction of Null Distribution for $n = 3$

Consider only the last three observations: 0.65, -0.40, 0.35.

The signed-rank transformation of this sample is: 3, -2, 1.

SR_1	SR_2	SR_3	\rightarrow	$\sum SR_i$
3	2	1	\rightarrow	6
3	2	-1	\rightarrow	4
3	-2	1	\rightarrow	2
3	-2	-1	\rightarrow	0
-3	2	1	\rightarrow	0
-3	2	-1	\rightarrow	-2
-3	-2	1	\rightarrow	-4
-3	-2	-1	\rightarrow	-6

Null Distribution

Calculating the P-Value

$$\begin{aligned} \text{p-value} &= P(T_{WSR} \geq T_{WSR}^*) \\ &= \#\{T_{WSR} \geq T_{WSR}^*\} / 2^n \\ &= 19/256 \\ &= 0.07421875 \end{aligned}$$

Alternative Null Distributions

Non-Data Distributions

- If no ties or zeros, null distribution can be computed without reference to data
- **Pre-computed tables** for small sample sizes
- **Recursive algorithms** for larger sample sizes

Approximate Distributions

- Same restrictions on ties and zeros
- **Asymptotic approximation** to normal distribution

$$\frac{T_{WSR}^*}{\sqrt{\frac{n(n+1)(2n+1)}{6}}} \xrightarrow{d} N(0, 1)$$

Comments

- 1 Popular by **historical accident**: alternative null distribution helpful when no computers
- 2 Asymptotic null distribution **saves computation cost** for large samples
 - 1 But with large samples, t test often works well anyway
- 3 **Ties and zero values** can be a headache unless null distribution is constructed by exhaustive permutation
- 4 Less sensitive to **outliers** than t test or Fisher Permutation test
- 5 Pretty good **asymptotic relative efficiency** compared to t test:
 - For normal distribution about $\frac{3}{\pi} = 95.5\%$

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Test Statistic

Focus on Median

Let $\mu \equiv 0$ be the hypothesized **median** value.

Sign Test Statistic

$$T_{ST} = \sum_{i=1}^n 1\{x_i > 0\}$$

x_i	0	0.2	1.75	1.6	-0.15	0.65	-0.4	0.35
$1\{x_i > 0\}$?	1	1	1	0	1	0	1

$$T_{ST}^* = 5$$

Deriving the Null Distribution

No Ties Requirement

If **no observations equal zero**, then any observation in the sample was equally likely to have been drawn with either a **positive** or **negative** sign.

Distribution of Signs

The sign of each observation is a random draw from a **Bernoulli** distribution with parameter $\theta = \frac{1}{2}$.

Null Distribution

The total number of positive signs (a sum of Bernoulli random variables) is **Binomial** distributed with parameters n and $\theta = \frac{1}{2}$.

Deriving the Null Distribution

Lower-Tail Test

$$\text{p-value} = \sum_{j=1}^{T_{ST}^*} \binom{n}{k} \left(\frac{1}{2}\right)^n$$

Upper-Tail Test

$$\text{p-value} = \sum_{j=1}^{n-T_{ST}^*} \binom{n}{k} \left(\frac{1}{2}\right)^n$$

Deriving the Null Distribution

The Problem with Ties to Zero

Three common ways to deal with zero-valued observations:

- 1 Assign ties conservatively ✓
- 2 Assign ties a value of $\frac{1}{2}$
- 3 Drop ties from sample ✓

Null Distribution

Calculating the Conservative P-value

Treat the zero observation as a **negative value**.

$$\begin{aligned} \text{p-value} &= \sum_{j=1}^3 \binom{8}{k} \left(\frac{1}{2}\right)^8 \\ &= 0.3632813 \end{aligned}$$

Calculating the Conditional P-value

Drop the zero-value, and treat as a sample with $n = 7$.

$$\begin{aligned} \text{p-value} &= \sum_{j=1}^2 \binom{7}{k} \left(\frac{1}{2}\right)^7 \\ &= 0.2265625 \end{aligned}$$

Comments

- 1 Concerns **median**, not (necessarily) **mean**
- 2 Without ties to zero, **no distributional requirements**
- 3 Valid even with **skewed** distributions
- 4 Generally **low efficiency**
- 5 Asymptotically relatively efficient for **very heavy-tailed** distributions

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Two-Sample Analogues

Matched-Pairs Samples

Take the **vector-difference** of the samples, then perform a one-sample test.

Independent Samples

There are two-sample analogues to each of the discussed tests:

- ① t test → two-sample t test
- ② Fisher Permutation test → Pitman Permutation test
- ③ Wilcoxon Signed-Rank test → Wilcoxon-Mann-Whitney Rank-Sum test
- ④ Sign test → Median test

Overview

Manski's Law of Decreasing Marginal Credibility

Tests based on stronger assumptions tend to yield **sharper** but **less credible** results.

Non-Parametric Tests

- 1 These are **not** “no assumption” or “distribution free” tests
- 2 **Relative efficiency** helps to decide between competing tests
- 3 No statistical test is **universally** appropriate